

Homework 4 Solutions.

(1)

It suffices to show that

$$\{\omega: X(\omega) \leq x\} = \{\omega: \omega \leq F(x)\} \quad (*)$$

Indeed, if this holds then

$$\begin{aligned} P\{\omega: X(\omega) \leq x\} &= P\{\omega: \omega \leq F(x)\} \\ &= P([0, F(x)]) = F(x) \end{aligned}$$

because P is Lebesgue measure, and we are done.

To check (*), we prove two inclusions:

" \supseteq ": Let $\omega \leq F(x)$. We want to show that $X(\omega) \leq x$.

This amounts to proving that for every y such that $F(y) < \omega$, we have $y \leq x$.

For every such y , $F(y) < \omega \leq F(x)$. Since F is nondecreasing, $y \leq x$ follows.

" \subseteq ": We argue by contraposition. Assume $X(\omega) \leq x$ but $\omega > F(x)$. Since F is right-continuous, $\exists \varepsilon > 0$ such that $F(x + \varepsilon) < \omega$. This means (by the definition of X) that $X(\omega) \geq x + \varepsilon > x$.

The contradiction completes the proof.

(3)

a) The density of a random variable is the derivative of its distribution function.

Let $F_{\phi(X)}$, F_X denote the distribution functions of $\phi(X)$ and X respectively.

Then

$$F_{\phi(X)}(z) = P(\phi(X) \leq z) = P(X \leq \phi^{-1}(z)) = F_X(\phi^{-1}(z)).$$

Therefore the density of $\phi(X)$ is

$$F'_{\phi(X)}(z) = F'_X(\phi^{-1}(z)) (\phi^{-1}(z))' = \boxed{\frac{f(\phi^{-1}(z))}{\phi'(\phi^{-1}(z))}}$$

b) $\phi^{-1}(x) = \frac{x-b}{a}$; $\phi'(x) = a$, hence

the density of $aX + b$ is ~~$\frac{1}{a} f(\frac{x-b}{a})$~~

$$\boxed{\frac{1}{a} f\left(\frac{x-b}{a}\right)}$$

c) $F_{X^2}(x) = P(X^2 \leq x) = P(-\sqrt{x} \leq X \leq \sqrt{x}) = F_X(\sqrt{x}) - F_X(-\sqrt{x})$.

Differentiating, we get

$$\frac{d}{dx} F_X(\sqrt{x}) = \frac{f(\sqrt{x})}{2\sqrt{x}}, \quad \frac{d}{dx} F_X(-\sqrt{x}) = -\frac{f(-\sqrt{x})}{2\sqrt{x}}$$

Thus the density of X^2 is

$$\boxed{\frac{f(\sqrt{x}) + f(-\sqrt{x})}{2\sqrt{x}}}$$

d) For $f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$, we get the density of X^2 equal to

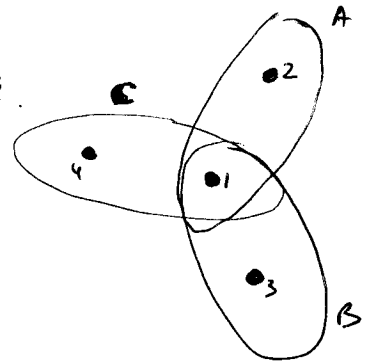
$$\boxed{\frac{1}{\sqrt{2\pi x}} e^{-x/2}} \leftarrow \chi^2\text{-distribution.}$$

(4)

Let $\Omega = \{1, 2, 3, 4\}$, $P =$ uniform distribution.

$A := \{1, 2\}$, $B := \{1, 3\}$, $C := \{1, 4\}$.

Then $P(A) = \frac{1}{2}$, $P(B) = \frac{1}{2}$, $P(C) = \frac{1}{2}$.



Now, $P(A \cap B) = \frac{1}{4} \Rightarrow A, B$ are independent

$P(B \cap C) = \frac{1}{4} \Rightarrow B, C$ independent

$P(A \cap C) = \frac{1}{4} \Rightarrow A, C$ independent

But $\left\{ \begin{array}{l} P(A \cup B) = \frac{3}{4} \\ P(C) = \frac{1}{2} \end{array} \right\}, P((A \cup B) \cap C) = \frac{1}{4}$

$\frac{3}{4} \cdot \frac{1}{2} \neq \frac{1}{4} \Rightarrow A \cup B$ is not independent of C .

A, B, C are not independent (otherwise, by a σ -algebra theorem, $A \cup B$ would be independent of C).

(5)

1) Let $|\Omega| = n$. The maximum number of events is 2^n (all subsets).
~~Moreover~~ Moreover we can make all 2^n subsets of Ω independent by defining a probability measure P that satisfies $P(\omega) = \frac{1}{2^n}$ for some point $\omega \in \Omega$.

(It is easy to check that all 2^n subsets are independent).

2) But there is more to this problem. Suppose we want to consider only probability measures for which

$$0 < P(A) < 1$$

for every proper $A \subset \Omega$.

Then there can be at most $\log_2 n$ independent events.

Let's prove this. Suppose A_1, \dots, A_m are independent and nonempty.

Then all sets of the form

$$A_1^{(c)} \cap A_2^{(c)} \cap A_3^{(c)} \cap \dots \cap A_m^{(c)},$$

where we put or don't put c (the complement)

for each A_k 's, are disjoint and non-empty.

(Indeed, every two such sets must differ in at least one place where one has A_k while the other has A_k^c ;

they must be disjoint. Also, by the independence,

each set has positive measure \Rightarrow non-empty.

~~Therefore~~ The number of such sets is 2^m (choose whether to put or not put c in 2^m ways).

Since they are disjoint, $n = |\Omega| \geq 2^m$.

Hence $m \leq \log_2 n$

(May be +1 for the empty set).

This bound is attained: let $\Omega = \{H, T\}^m$
 where $m = \log_2 n$ for some n . Then $|\Omega| = n$.

Let $A_k = \{H \text{ in } k\text{th toss of a coin}\}$, $k = 1, \dots, m$

These are independent.

(6)

$$P(X < y+x | X \geq y) = \frac{P(y \leq X < y+x)}{P(X \geq y)} = \frac{F(x+y) - F(y)}{1 - F(y)}$$

By the assumption, the LHS = $F(x)$.

Let $\Phi(x) = 1 - F(x)$; then we get

$$\frac{\Phi(y) - \Phi(x+y)}{\Phi(y)} = 1 - \Phi(x)$$

from which it follows that

$$\Phi(x+y) = \Phi(x) \cdot \Phi(y)$$

Taking logarithms we see that

$$(\log \Phi)(x+y) = (\log \Phi)(x) + (\log \Phi)(y) \quad \forall x, y \geq 0$$

which means that $\log \Phi$ is a linear function:

$$(\log \Phi)(x) = ax + b$$

Hence $\Phi(x) = e^{ax+b}$, $x \geq 0$

so $F(x) = 1 - e^{ax+b}$

The condition ~~$F(0) = 0$~~ $X \geq 0$ means $F(0) = 0 \Rightarrow b = 0$

The condition $F(x) \rightarrow 1$ as $x \rightarrow \infty$ means $a < 0 \Rightarrow$

$$a = -\lambda \text{ for some } \lambda > 0.$$

Therefore $F(x) = 1 - e^{-\lambda x}$