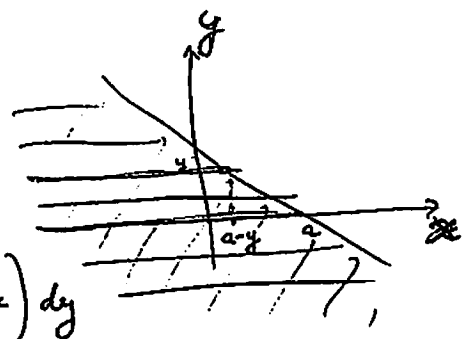


• Sums of general indep. r.v.'s.  $X, Y$ .

PDF:  $F_{X+Y}(a) = P\{X+Y \leq a\} = \iint_{(X,Y) \in \mathbb{R}^2: X+Y \leq a} f_{X,Y}(x,y) dx dy = \iint f_X(x) f_Y(y) dx dy$  by indep



$$= \int_{-\infty}^{\infty} \int_{-\infty}^{a-y} f_X(x) f_Y(y) dx dy = \int_{-\infty}^{\infty} f_Y(y) \left( \int_{-\infty}^{a-y} f_X(x) dx \right) dy$$

$\underbrace{\int_{-\infty}^{a-y} f_X(x) dx}_{F_X(a-y)}$

~~$$\int_{-\infty}^{\infty} \int_{-\infty}^{a-y} f_X(x) f_Y(y) dx dy$$~~

$$= \int_{-\infty}^{\infty} F_X(a-y) f_Y(y) dy.$$

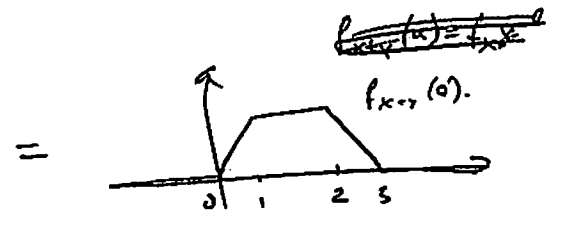
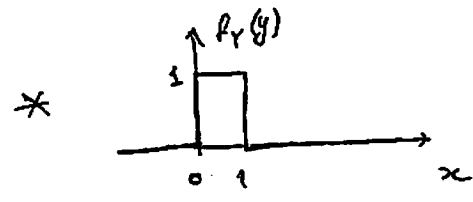
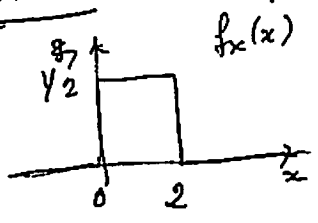
Differentiate w.r to a.

PDF:  $f_{X+Y}(a) = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy = (f_X * f_Y)(a).$

PROP  $X, Y$  indep. continuous.  $\Rightarrow$  the PDF of  $X+Y$  is given by the convolution of PDF's of  $X, Y$ ;

$$f_{X+Y}(a) = (f_X * f_Y)(a) = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy.$$

• Rework Ex. on p. 59 in terms of convolutions!

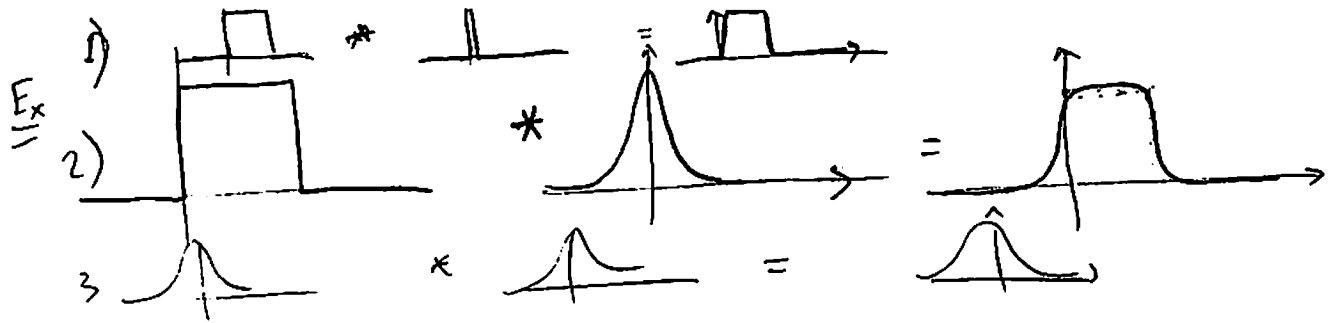


~~Study Ex 5.1~~

Check for  $a < 1$ :  $f_{X+Y}(a) = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy$

$\underbrace{\int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy}_{\substack{\text{for } a < 1, \text{ } a-y \in (0,2), \text{ } y \in (0,1) \\ \text{so } f_X(a-y) = 1/2, \text{ } f_Y(y) = 1 \\ \text{so } \int_0^a 1/2 \cdot 1 dy = a/2}}$

$$= \int_0^a \frac{1}{2} dy = a/2.$$



• Sums of independent normals is normal

Ex  ~~$X \sim N(0, \sigma^2)$~~   $X, Y \sim N(0, 1)$  indep.  $X+Y \sim ?$

$$f(a) = \int_{-\infty}^{\infty} f(a-y) f(y) dy = \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-(a-y)^2/2} e^{-y^2/2} dy$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-(y^2 - ay + a^2/2)} dy$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-[(y - \frac{a}{2})^2 + \frac{a^2}{4}]} dy = e^{-a^2/4} \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-(y - \frac{a}{2})^2} dy$$

$$\begin{aligned} (y - \frac{a}{2})^2 &= \frac{x^2}{2} \\ y - \frac{a}{2} &= \frac{x}{\sqrt{2}} \end{aligned} \quad e^{-a^2/4} \frac{1}{2\sqrt{\pi}} \int_{-\infty}^{\infty} e^{-x^2/2} dx = \left( \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-x^2/2} dx \right) \cdot \frac{1}{2\sqrt{\pi}} e^{-a^2/4}$$

$$= \frac{1}{\sqrt{2\pi}} e^{-a^2/4} = \frac{1}{\sqrt{2\pi} \sigma} e^{-a^2/2\sigma^2} \text{ for } \sigma = \sqrt{2}$$

$$\boxed{X+Y \sim N(0, 2)}$$

More generally:

THM  $X_i \sim N(\mu_i, \sigma_i)$  indep.  $\Rightarrow \sum_{i=1}^n X_i \sim N(\mu, \sigma)$  where

$$\begin{aligned} \mu &= \sum_{i=1}^n \mu_i & \sigma^2 &= \sum_{i=1}^n \sigma_i^2 \\ \uparrow & & \uparrow & \\ \text{(mean)} & & \text{(variance)} & \end{aligned}$$

In particular, if  $\mu_1 = \mu_2 = 0$ ,  $\sigma_1 = \sigma_2 = 1 \Rightarrow \sigma = \sqrt{2}$ .

Ex [Ghahramani Ex 11.1] Class = 25 students

Students' grades in an exam are normally distributed with mean 72 and variance 15  
 Prob { average grade of this ~~course~~ ~~class~~ of  $\geq 74$  } = ?

$X_1, \dots, X_n = \text{grades}, X_k \sim N(\mu, \sigma^2), \mu = 72, \sigma^2 = 15$

Ave.  $\bar{X} = \frac{1}{n} \sum_{k=1}^n X_k \sim N\left(\frac{1}{n} \sum \mu, \frac{1}{n} \sum \sigma^2\right)$  by THM p. 67

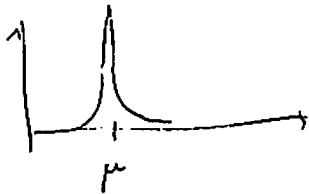
$= N\left(\mu, \frac{\sigma^2}{n}\right)$

$= N(72, 0.6)$

$1 - \Phi(2.58) = 0.0049$

$P\{\bar{X} > 74\} = P\left\{\frac{\bar{X} - 72}{\sqrt{0.6}} > \frac{74 - 72}{\sqrt{0.6}}\right\} = P\{\bar{X} - 72 > 1 - \Phi(3) = 0.0013\}$

Remark on  $N\left(\mu, \frac{\sigma^2}{n}\right)$ : Variance small  $\Rightarrow$  the bigger the sample, the better the higher the estimate of  $\mu$  by  $\bar{X}$



Sums of independent discrete r.v.'s.

- THM. 1. Let  $X_1, \dots, X_n \sim$  ~~Bernoulli~~ Bernoulli( $p$ ) indep.  
 Then  $X_1 + \dots + X_n \sim \text{Binom}(n, p)$
2. Let  $X \sim \text{Binom}(n, p), Y \sim \text{Binom}(m, p)$ . Then  $X + Y \sim \text{Binom}(n+m, p)$ .
3. Let  $X \sim \text{Poisson}(\lambda), Y \sim \text{Poisson}(\mu)$ . Then  $X + Y \sim \text{Poisson}(\lambda + \mu)$ .

1, 2 - explain in terms of successes/trials.

3: <sup>Express</sup> ~~let~~  $\lambda = np, \mu = mp$  where  $n, m \rightarrow \infty, p \rightarrow 0$ . Then, ~~asympt~~ in the limit,  
 $X \sim \text{Binom}(n, p), Y \sim \text{Binom}(m, p)$ .  
 By part 2,  $X + Y \sim \text{Binom}(n+m, p) \approx \text{Poisson}((n+m)p)$  in the limit  
 $= \text{Poisson}(\lambda + \mu)$