

## Remarks

1. If  $f$  is a density then

- $f \geq 0$  pointwise

(~~QED~~ since  $f = F' \geq 0$  as  $F$  is  $\uparrow$ ).

- $\int_{-\infty}^{\infty} f(x) dx = 1$

( $= \lim_{a \rightarrow \infty} F(a) = 1$ )

- $\int_a^b f(x) dx = F(b) - F(a)$

2. If  $X$  is abs. continuous then

$$P\{X=x\} = 0 \quad \forall x \in \mathbb{R}$$

( $= F(x) - F(x-) = 0$  by continuity of  $F$ )

3.  $\exists$  r.v's that are neither discrete nor continuous. (Why?)

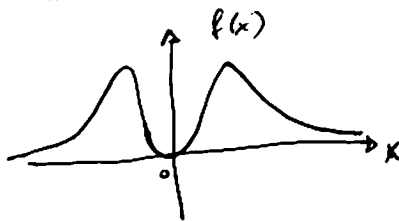
## Functions of r.v's: Change of density

### Ex. 1. (Range of density)

Ex. 1. Let  $X \sim N(0,1)$ ;  $Y := X^3$ . density of  $Y = ?$

CDF:  $F_Y(a) = P\{Y \leq a\} = P\{X^3 \leq a\} = P\{X \leq a^{1/3}\} = F_X(a^{1/3})$

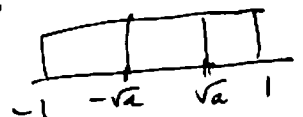
Density:  $f_Y(a) := \frac{d}{da} F_Y(a) = f_X(a^{1/3}) \cdot \frac{1}{3} |a|^{-2/3} = \frac{1}{3|a|^{2/3} \sqrt{2\pi}} e^{-\frac{|a|^{2/3}}{2}}, a \neq 0$



Ex. 2 Let  $X \sim \text{Unif}[-1,1]$ ,  $Y = X^2$ .

CDF:  $F_Y(a) = P\{Y \leq a\} = P\{X^2 \leq a\} = P\{|X| \leq \sqrt{a}\} = \sqrt{a}, 0 \leq a \leq 1$

$f_Y(a) = \frac{d}{da} F_Y(a) = \begin{cases} \frac{1}{2\sqrt{a}}, & 0 < a \leq 1 \\ 0, & - \end{cases}$



## 2.5. Expectations of discrete r.v's.

Def. Let  $X$  be a discrete r.v. with PMF  $p(x_k) = P\{X=x_k\}$ .  
The expectation (expected value) of  $X$  is defined as

$$E[X] = \sum_k x_k p(x_k).$$

We require that this series converges absolutely;  
otherwise we say that  $X$  has no expectation.

Ex.  $X = \#$  heads in 3 coin flips:

$$E[X] = 0 \cdot \frac{1}{8} + 1 \cdot \frac{3}{8} + 2 \cdot \frac{3}{8} + 3 \cdot \frac{1}{8} = \boxed{1.5}$$

Remark If all  $N$  values of  $X$  are equally likely then  $p(x_k) = \frac{1}{N} \Rightarrow$   
 $E[X] = \frac{1}{N} \sum_{k=1}^N x_k$  (arithmetic mean)

Otherwise, weighted by  $p(x_k)$ .

Ex (lottery) 6 out of 49. ~~Prizes:~~

| # correct | Prize       |
|-----------|-------------|
| all 6     | \$1,200,000 |
| 5         | \$800       |
| 4         | \$35        |

Expected amount the player wins?

| correct | not |
|---------|-----|
| 6       | 43  |

PMF:  $p(1,200,000) = \frac{1}{\binom{49}{6}}$

$$p(800) = P\{5 \text{ correct, 1 not}\} = \frac{\binom{6}{5} \binom{43}{1}}{\binom{49}{6}}$$

$$p(35) = P\{4 \text{ correct, 2 not}\} = \frac{\binom{6}{4} \binom{43}{2}}{\binom{49}{6}}$$

$$E[X] = 1,200,000 \cdot p(1,200,000) + 800 \cdot p(800) + 35 \cdot p(35) + 0 \cdot p(0) = \boxed{0.13}$$

Ans:  $\boxed{134}$

# Ex (Repeated values) (~ Remark 2.26)



Suppose  $\Omega = \bigcup_k \Omega_k$  be a <sup>countable</sup> partition of a sample space

Assume that  $X$  is constant on each  $\Omega_k$ , say  $X = x_k$  on  $\Omega_k$ .

Then  $E[X] = \sum_k x_k P(\Omega_k)$

possibly the same  $x_k$  for different  $\Omega_k$

If  $X$  takes different values on each  $\Omega_k \Rightarrow$  QED by def of  $E(X)$   
 If not, Combine the parts  $\Omega_k$  on which  $X$  takes the same value into a bigger set.

Prop (Linearity).  $\forall$  r.v.  $X, Y, \forall$  const  $a \in \mathbb{R}$ :

$$(i) E[X+Y] = E[X] + E[Y]$$

$$(ii) E[aX] = a \cdot E[X]$$

(i)  $\exists$  partition  $\Omega = \bigcup_{ij} \Omega_{ij}$  s.t. both  $X, Y$  are const on each  $\Omega_{ij}$  (Ex.)  
 $\Omega_{ij} = \{X=x_i, Y=y_j\}$

$$\begin{aligned} \Rightarrow E[X+Y] &= \sum_{ij} (x_i + y_j) P(\Omega_{ij}) \\ &= \sum_i x_i \underbrace{\sum_j P(\Omega_{ij})}_{P\{X=x_i\}} + \sum_j y_j \underbrace{\sum_i P(\Omega_{ij})}_{P\{Y=y_j\}} \\ &= E[X] + E[Y] \quad (\text{by def.}) \end{aligned}$$

Remark: (1) Absolute continuity was used here.  
 (2) For more than 2:

$$E\left[\sum_i X_i\right] = \sum_i E[X_i]$$

Other properties (see book):

- $X \leq Y \Rightarrow E[X] \leq E[Y]$
- $X = c \text{ (const)} \Rightarrow E[X] = c$
- $a < X \leq b \Rightarrow a \leq E[X] \leq b$
- $f: \mathbb{R} \rightarrow \mathbb{R} \Rightarrow E[f(X)] = \sum f(x_i) P\{X_i = x_i\}$   
 (follows from Ex.)
- $E[1_E] = P(E)$

### Ex (Matching Problem)

$N$  hw returned to  $N$  students at random.

$$E(\# \text{ students getting their own hw}) = ?$$

(Recall that we calculated ~~that at least 1 stud~~  
 $P\{X \geq 1\} = 1 - \frac{1}{e}$ ).

Method of indicators:

~~$X = \# \text{ students}$~~   $X = \sum_{i=1}^N X_i$  where  $X_i = \begin{cases} 1, & \text{student } i \text{ gets own hw} \\ 0, & \text{ } \end{cases} = \mathbb{1}_{\{i \text{ gets own hw}\}}$

$$\Rightarrow E[X] = \sum_{i=1}^N E[X_i].$$

$$E[X_i] = \cancel{0.5} P\{\text{Student } i \text{ gets own hw}\} = \frac{1}{N}.$$

$$\Rightarrow E[X] = N \cdot \frac{1}{N} = \boxed{1}.$$

Property: If  $X_1, X_2, \dots$  are random variables, then

$$E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i]$$

(To be proved later)

### Example (Group Testing)

Need to test blood of  $n$  people for a rare disease (syphilis in men drafted during WW II)  
 $p = P(\text{positive})$ .

Method 1: test all people individually ( $\Rightarrow n$  tests)

Method 2: draw and mix blood of  $k$  people.

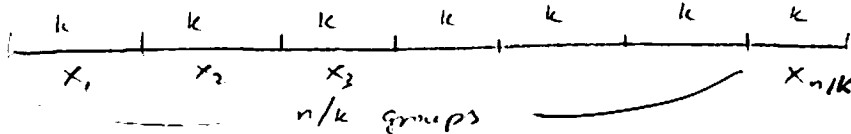
(Group testing) If test is negative  $\Rightarrow$  no more tests needed ( $\Rightarrow 1$  test)

If positive  $\Rightarrow$  test all  $k$  people individually ( $\Rightarrow k+1$  tests).

Repeat for next  $k$  people, etc.

Expected # of tests in Method 2?

Divide  $n$  people into  $n/k$  groups of  $k$ :



$X_i := \#(\text{tests on group } i)$ .



$$X = X_1 + X_2 + \dots + X_{n/k}$$

~~Value of tests needed for group~~  $X_i$  takes values 1,  $k+1$ .

$$P\{X_i = 1\} = P\{\text{all } k \text{ people are negative}\} = (1-p)^k$$

$$P\{X_i = k+1\} = 1 - (1-p)^k$$

$$E[X_i] = 1 \cdot (1-p)^k + (k+1) [1 - (1-p)^k] \quad (\text{same for each group } i)$$

$$\Rightarrow E[X] = \sum_{i=1}^{n/k} E[X_i] = \frac{n}{k} \left[ (1-p)^k + (k+1) [1 - (1-p)^k] \right]$$

For example, if  $n = 100,000$ ,  $p = 10^{-4}$  (on ave, 10 sick)

With  $k = 120 \Rightarrow E[X] = 2026$

Compare to method 1 where  $\# = 100,000$ .