

(02/02)

3.5. Standard Normal distribution (cont.)

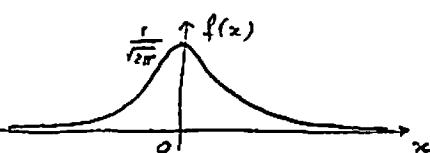
(02/02/2020)

Def A rv X has the standard normal distribution if the pdf of X is

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}, \quad -\infty < x < \infty.$$

Notation:

$$X \sim N(0, 1).$$



- Gauss used normal dist to model his observations in astronomy so "normal" is often called "Gaussian".

- Why $\frac{1}{\sqrt{2\pi}}$?

Prop $f(x)$ above is indeed a pdf, i.e.

$$\int_{-\infty}^{\infty} f(x) dx = 1$$

Proof Need to show:

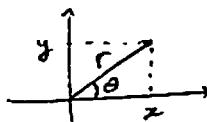
$$I := \int_{-\infty}^{\infty} e^{-x^2/2} dx = \sqrt{2\pi}$$

$$\text{Trick: } I^2 = \left(\int_{-\infty}^{\infty} e^{-x^2/2} dx \right) \left(\int_{-\infty}^{\infty} e^{-y^2/2} dy \right) = \int_{\mathbb{R}^2} e^{-(x^2+y^2)/2} dx dy \quad (\text{by Fubini Thm}).$$

Pass to polar coordinates:

$$x = r \cos \theta, \quad y = r \sin \theta$$

$$dx dy = r dr d\theta.$$



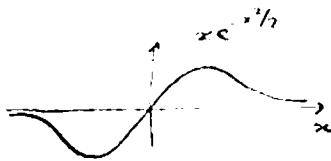
$$I^2 = \int_0^{2\pi} \int_0^{\infty} e^{-r^2/2} r dr d\theta = 2\pi \int_0^{\infty} r e^{-r^2/2} dr \stackrel{\begin{array}{l} u = r^2/2 \\ du = r dr \end{array}}{=} 2\pi \int_0^{\infty} e^{-u} du = 2\pi.$$

Hence $I = \sqrt{2\pi}$. QED.

Prop For $X \sim N(0,1)$, $E[X] = 0$ and $\text{Var}(X) = 1$ (Thus $\sigma_X = 1$)

Proof $E[X] = \frac{1}{\sqrt{2\pi}} \int_0^\infty x e^{-x^2/2} dx = 0$

odd function



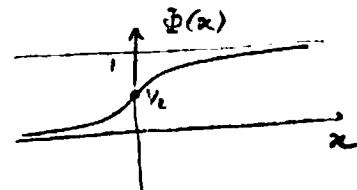
$$\text{Var}(X) = E[X^2] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^\infty x^2 e^{-x^2/2} dx \quad \textcircled{=}$$

(By parts $u = x$, $dv = x e^{-x^2/2} dx \Rightarrow v = \int x e^{-x^2/2} dx = -e^{-x^2/2}$)

$$\textcircled{=} \frac{1}{\sqrt{2\pi}} \left[-x e^{-x^2/2} \right]_{-\infty}^\infty + \int_{-\infty}^\infty e^{-x^2/2} dx = 1 \quad \text{by Prop. p. 48. O.F.D.}$$

• Cdf of $N(0,1)$ is denoted $\Phi(a)$, recall $\Phi(a) = P\{X \leq a\}$

$$\boxed{\Phi(a) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^a e^{-x^2/2} dx, \quad x \in \mathbb{R}.}$$



$\Phi(x)$ can not be expressed in terms of elementary functions like $\sin x$, e^x , etc.

It is a "special function", tabulated on p 201 of Ross

On a computer, can be computed as $\Phi(x) = \frac{1}{2} + \frac{1}{2} \text{erf}\left(\frac{x}{\sqrt{2}}\right)$

"erf function", another special function

Ex Let $X \sim N(0,1)$. Compute $P\{|X| \leq 2\}$

$$P\{-2 \leq X \leq 2\} = \Phi(2) - \Phi(-2).$$

By the symmetry of the pdf $f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$ (even function), we have

$$\Phi(-2) = 1 - \Phi(2)$$

$$P\{|X| \leq 2\} = 2\Phi(2) - 1 = 0.954.$$

⇒ fast tail decay of $f(x)$

$$P\{|X| \leq 3\} = 0.997$$

• Remark: By the same symmetry reasoning,

$$\boxed{\Phi(-x) = 1 - \Phi(x)}$$

for all x

Ex Let $X \sim N(0, 1)$ and $Y = \frac{\mu}{\mu + \sigma X}$ ($\mu \in \mathbb{R}, \sigma > 0$)
 Find the pdf, mean, variance of Y .

CDF: $F_Y(x) = P\left\{ \frac{\mu}{\mu + \sigma X} \leq x \right\} = P\left\{ X \leq \frac{x-\mu}{\sigma} \right\} = F_X\left(\frac{x-\mu}{\sigma}\right).$

PDF: $f_Y(x) = \frac{d}{dx} F_Y(x) = \frac{1}{\sigma} f_X\left(\frac{x-\mu}{\sigma}\right) = \frac{1}{\sigma\sqrt{2\pi}} e^{-(x-\mu)^2/\sigma^2}.$

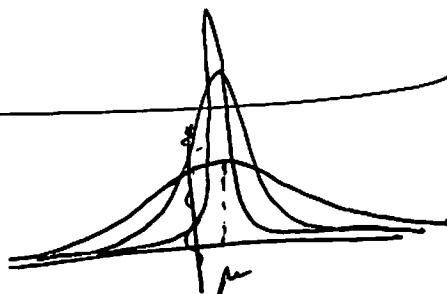
Def Mean: $E[Y] = E[\mu + \sigma X] = \mu$

Def Var: $Var(Y) = \sigma^2 Var(X) = \sigma^2$.

Def (General normal distr) $\equiv X \sim N(\mu, \sigma^2)$ if X has density

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-(x-\mu)^2/\sigma^2}.$$

1. X has mean μ and variance σ^2 .

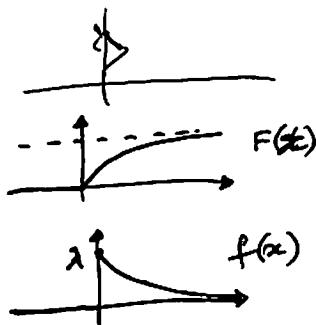


Exponential distribution

Def $X \sim Exp(\lambda)$ if $P(X > t) = e^{-\lambda t}$, $t \geq 0$.
 "rate"

$$\Rightarrow \text{CDF: } F(x) = 1 - e^{-\lambda x}$$

$$\text{PDF: } f(x) = F'(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & x < 0 \end{cases}$$



$$E[X] = \int_{-\infty}^{\infty} x f(x) dx$$

$$= \int_0^{\infty} \lambda x e^{-\lambda x} dx = \frac{1}{\lambda} \int_0^{\infty} y e^{-y} dy = \dots (\text{by parts}) = \left(\frac{1}{\lambda}\right).$$

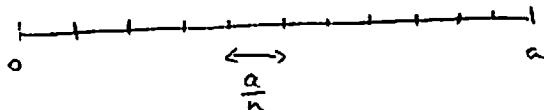
$$E[X^2] = \int_{-\infty}^{\infty} x^2 f(x) dx = \dots = \frac{2}{\lambda^2}. \quad \Rightarrow \quad \text{Var}(x) = \left(\frac{2}{\lambda^2}\right) - \left(\frac{1}{\lambda}\right)^2 = \left(\frac{1}{\lambda^2}\right).$$

WAITING TIME

- Exponential distr is a typical model for waiting time (continuous) until some event occurs. WHY?

Ex | $X = \text{time of the first call at a police station (say, after midnight).}$
Let's find the distribution of X .

$P\{X \leq a\} = ?$ Divide $[0, a]$ into n intervals of length $\frac{a}{n}$ (n large; e.g. every interval is a second)



$P\{\text{call during a given interval}\} = \lambda \cdot \frac{a}{n}$, for some coeff. λ (this prob. is proportional to the length of the interval)

$$\begin{aligned} \Rightarrow P\{X > a\} &= P\{\text{no call in } [0, a]\} = P\{\text{no call during each interval}\} \\ &= \left(1 - \frac{\lambda a}{n}\right)^n \quad (\text{by independence}) \\ &\rightarrow e^{-\lambda a} \quad \text{as } n \rightarrow \infty \end{aligned}$$

Hence $\boxed{X \sim \text{Exp}(\lambda)}$

• Prop (Memoryless Property) (let $X \sim \text{Exp}(\lambda)$) Then

$$P\{X > t+s \mid X > t\} = P\{X > s\} \quad \text{for } s, t > 0$$

\uparrow $P\{\text{need to wait } > s \text{ more minutes}\}$ \uparrow $P\{\text{need to wait } > s \text{ minutes}\}$
after having waited t min

Proof: L.H.S = $\frac{P\{X > t+s\}}{P\{X > t\}} = \frac{e^{-\lambda(t+s)}}{e^{-\lambda t}} = e^{-\lambda s} = \text{R.H.S.} \quad \text{Q.E.D.}$

THM Exponential distribution is the only continuous memoryless distribution

Applications. All of these have exponential distribution.

- time until next email message arrives
- time until next customer enters the store
- lifetime of equipment (e.g. a cellphone) if there is no aging;
can only die due to an accident
- time between successive independent events (accidents, e-mails, payments, etc.)

Exponential distr. is a "continuous version" of geometric distr.

Prop Let $X \sim \text{Exp}(\lambda)$. Then $\lceil X \rceil \sim \text{Geom}(p)$, where $p = 1 - e^{-\lambda}$.

$$\begin{aligned} P\{\lceil X \rceil = k\} &= P\{k-1 < X \leq k\} = P\{X > k-1\} - P\{X > k\} \\ &= e^{-\lambda(k-1)} - e^{-\lambda k} = e^{-\lambda(k-1)} \underbrace{(1 - e^{-\lambda})}_{p} = (1-p)^{k-1} \cdot p. \end{aligned}$$

Ex A small store is visited by 3 customers per hour on average. What is the prob that the first two customers will enter within 5 min from each other?

Memoryless \Rightarrow start the clock after 1st customer enters.

$X = \text{time until next customer} \sim \text{Exp}(3)$.

$$P\{X < \frac{1}{12}\} = 1 - e^{-3 \cdot \frac{1}{12}} = \underbrace{0.22}_{\text{5 min}},$$

Cauchy distribution

$X \sim \text{Cauchy}$ if X has density

$$f(x) = \frac{1}{\pi(1+x^2)}$$



Ex (Check $\int_{-\infty}^{\infty} f(x) dx = 1$).

X does NOT have expectation:

$$E[X] = \frac{1}{\pi} \int_{-\infty}^{\infty} \frac{(x)}{\pi(1+x^2)} dx = +\infty.$$

$\approx \frac{1}{x}$ as $x \rightarrow \infty$

Example: ~~Intensity of light~~

Intensity of light:

\Rightarrow Cauchy

