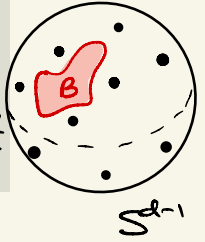


Example (Probabilistic Method)

Mark n dots on a soccer ball.

Let B be a sticker whose area is $< \frac{1}{n}$ of the ball's surface area.

Then B can be placed on the ball so it does not cover any dot.



Proof Let the dots be $x_1, \dots, x_n \in S^{d-1}$.

Choose $U \in SO(n)$ at random, according to Haar measure

$$\bullet \forall i: \underbrace{P\{x_i \in U(B)\}}_{\parallel E_i} = \underbrace{P\{U^{-1}(x) \in B\}}_S = \frac{\text{Area}(B)}{\text{Area}(S^{d-1})} < \frac{1}{n}$$

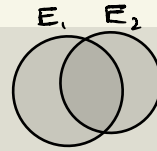
$\text{Unif}(S^{d-1})$ [uniqueness of rotation-invariant prob. measure on S^{d-1}]

• Union Bound:

$$P\{\exists i \in \{1, \dots, n\}: x_i \in U(B)\} = P\left(\bigcup_{i=1}^n E_i\right) \leq \sum_{i=1}^n P(E_i) < n \cdot \frac{1}{n} = 1.$$

$$\Rightarrow P\{\forall i \in \{1, \dots, n\}: x_i \notin U(B)\} \Rightarrow \exists \text{ such } U. \quad \square$$

Lem (Inclusion-Exclusion Principle)



$$\bullet P(E_1 \cup E_2) = P(E_1) + P(E_2) - P(E_1 \cap E_2)$$

$$\bullet P(E_1 \cup E_2 \cup E_3) = P(E_1) + P(E_2) + P(E_3) - P(E_1 \cap E_2) - P(E_1 \cap E_3) - P(E_2 \cap E_3) + P(E_1 \cap E_2 \cap E_3)$$

$$\bullet P\left(\bigcup_{i=1}^n E_i\right) = \sum_{i=1}^n P(E_i) - \sum_{i < j} P(E_i \cap E_j) + \sum_{i < j < k} P(E_i \cap E_j \cap E_k) - \dots + (-1)^{n-1} P(E_1 \cap \dots \cap E_n)$$

\uparrow
n terms

\uparrow
 $\binom{n}{2}$ terms

\uparrow
 $\binom{n}{3}$ terms

\uparrow
 $\binom{n}{n} = 1$ term

Proof for $n=2$: $E := E_1 \cup E_2 \Rightarrow (\mathbb{1}_E - \mathbb{1}_{E_1})(\mathbb{1}_E - \mathbb{1}_{E_2})(\omega) = 0 \quad \forall \omega \in \Omega$

$$\Rightarrow (\mathbb{1}_{E_1 \cap E_2} - \mathbb{1}_{E_1 \cap E_1} - \mathbb{1}_{E_1 \cap E_2} + \mathbb{1}_{E_1 \cap E_2})(\omega) = 0$$

$$\Rightarrow (\mathbb{1}_E - \mathbb{1}_{E_1} - \mathbb{1}_{E_2} + \mathbb{1}_{E_1 \cap E_2})(\omega) = 0$$

Rearrange terms $\Rightarrow \mathbb{1}_E(\omega) = (\mathbb{1}_{E_1} + \mathbb{1}_{E_2} - \mathbb{1}_{E_1 \cap E_2})(\omega) \quad \forall \omega \in \Omega$

Integrate $\Rightarrow \int \mathbb{1}_E dP = \int \mathbb{1}_{E_1} dP + \int \mathbb{1}_{E_2} dP - \int \mathbb{1}_{E_1 \cap E_2} dP$
 $\parallel \quad \parallel \quad \parallel \quad \parallel$
 $P(E) \quad P(E_1) \quad P(E_2) \quad P(E_1 \cap E_2)$

HW: prove for $\forall n$

Q. (derangements) n exams are returned to n students at random. What is the probability that no student receives their own exam?

$\Leftrightarrow P(\text{a random permutation } \pi \in S_n \text{ has no fixed point}) = ?$
↑
"derangement"

Sol: Complement: $E :=$ "at least one student gets own exam"

$$E = \bigcup_{i=1}^n E_i, \text{ where } E_i = \text{"student } i \text{ gets own exam"}$$

Use IEP:

$$(1) P(E_i) = \frac{|E_i|}{|S|} = \frac{(n-1)!}{n!} = \frac{1}{n}$$

← # of ways to return $n-1$ exams (after student i gets their own)

$$(2) P(E_i \cap E_j) = P\{\text{both students } i, j \text{ get their own exams}\}$$

$$= \frac{|E_i \cap E_j|}{|S|} = \frac{(n-2)!}{n!}$$

← # ways to return $n-2$ exams (after studs i, j get their own)

$$(3) P(E_i \cap E_j \cap E_k) = \frac{(n-3)!}{n!}$$

... IEP \Rightarrow

$$P(E) = n \cdot \frac{1}{n} - \binom{n}{2} \frac{(n-2)!}{n!} + \binom{n}{3} \frac{(n-3)!}{n!} - \dots + (-1)^{n-1} \binom{n}{n} \frac{(n-n)!}{n!}$$

$\frac{n!}{2!(n-2)!}$ $\frac{n!}{3!(n-3)!}$

$$= 1 - \frac{1}{2!} + \frac{1}{3!} - \dots + (-1)^{n-1} \frac{1}{n!}$$

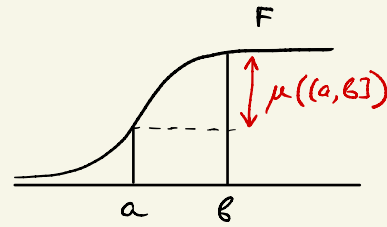
$$\Rightarrow P(E^c) = 1 - P(E) = \frac{1}{2!} - \frac{1}{3!} + \dots + \frac{(-1)^n}{n!} \approx \boxed{\frac{1}{e}} \approx 0.37$$

(recall $e^x = \sum_{n=0}^{\infty} \frac{x^n}{n!}$ $e^{-1} = \sum_{n=0}^{\infty} \frac{(-1)^n}{n!}$)

Describing a probability measure with a distribution function

Recall the def of Lebesgue-Stieltjes measure :

Thm \forall nondecreasing, right-continuous function $F: \mathbb{R} \rightarrow \mathbb{R}$
 $\exists!$ Borel measure μ on \mathbb{R} s.t.
 $\mu((a, b]) = F(b) - F(a) \quad \forall a \leq b$



Ex $F(x) = x \Rightarrow$ Lebesgue measure.

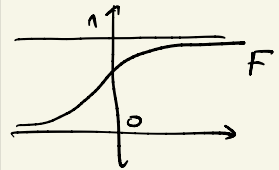
Note : If $\lim_{x \rightarrow -\infty} F(x) = 0$, $\lim_{x \rightarrow \infty} F(x) = 1$ then μ is a probability measure

$\left. \begin{array}{l} (-n, n] \uparrow \mathbb{R} \Rightarrow \text{by continuity, } \mu((-n, n]) \uparrow \mu(\mathbb{R}) \\ \parallel \\ F(n) - F(-n) \rightarrow 1 - 0 = 1 \end{array} \right\} \Rightarrow \mu(\mathbb{R}) = 1$

This justifies:

Def A distribution function is a function $F: \mathbb{R} \rightarrow [0, 1]$ such that :

- (i) F is nondecreasing;
- (ii) F is right-continuous;
- (iii) $\lim_{x \rightarrow -\infty} F(x) = 0$, $\lim_{x \rightarrow +\infty} F(x) = 1$.



We proved :

Thm \forall distribution function $F \exists!$ Borel probability measure μ on \mathbb{R}
such that
 $\mu((a, b]) = F(b) - F(a) \quad \forall a \leq b$.

Ex (a) if $F = \begin{cases} 0 & x < 2 \\ 1/2 & 2 \leq x < 3 \\ 1 & x \geq 3 \end{cases}$ then $\mu(\{2\}) = \mu(\{3\}) = \frac{1}{2}$.

(b) if $F = \begin{cases} 0 & x < 0 \\ x & 0 \leq x < 1 \\ 1 & x \geq 1 \end{cases}$ then $\mu =$ Lebesgue on $[0, 1]$.

Conversely:

Thm \forall Borel probability measure $\mu \exists!$ distribution function F such that
 $\mu(a, b] = F(b) - F(a) \quad \forall a \leq b.$

Proof Define $F(x) := \mu(-\infty, x], \quad x \in \mathbb{R}$

Then: (i) F is nondecreasing (monotonicity of measures)

(ii) F is right-continuous

$\lceil \forall y_n \downarrow x, (-\infty, y_n] \downarrow (-\infty, x]$
 $\Rightarrow \mu(-\infty, y_n] \downarrow \mu(-\infty, x]$ (continuity of measures)
 $\Rightarrow F(y_n) \downarrow F(x).$ \lrcorner

(iii) $\lim_{x \rightarrow -\infty} F(x) = 0, \quad \lim_{x \rightarrow +\infty} F(x) = 1$

\lceil let's prove this \nearrow

If $x_n \rightarrow +\infty, (-\infty, x_n] \uparrow \mathbb{R}$

$\Rightarrow \mu(-\infty, x_n] \uparrow \mu(\mathbb{R}) = 1$ (continuity of measures)

$\Rightarrow F(x_n) \uparrow 1.$ \lrcorner

QED