

EXPECTATION OF A RANDOM VARIABLE

Recall def of Lebesgue integral (measure theory, Durrett 1.4):

$(\Omega, \mathcal{F}, \mu)$ prob. space, $f: \Omega \rightarrow \mathbb{R}$ measurable.

- For indicator functions $\int \mathbb{1}_E d\mu = \mu(E)$
- For simple functions $\int (\sum_i a_i \mathbb{1}_{E_i}) d\mu = \sum_i a_i \mu(E_i)$
↑ disjoint
- For nonnegative mble functions

$$\int f d\mu := \sup \left\{ \int g d\mu : 0 \leq g \leq f, g \text{ mble} \right\}$$

• For \forall mble functions: $f = f^+ + f^-$, $\int f d\mu := \int f^+ d\mu - \int f^- d\mu$.

f is integrable $\iff \int f d\mu$ exists $\in \mathbb{R}$ (not $\pm\infty$) $\iff f^+, f^-$ integrable $\iff |f|$ integrable.

$$\int_E f d\mu := \int f \mathbb{1}_E d\mu.$$

Def Let X be a r.v. on (Ω, \mathcal{F}, P) . Then

$$E[X] := \int X dP$$

"expected value", "mean", "expectation"

Ex $X = \#$ K's in 2 flips

Prob: $\frac{1}{4} \quad \frac{1}{4} \quad \frac{1}{4} \quad \frac{1}{4}$
 $\Omega = \{KK, KT, TK, TT\}$
 $X \downarrow \quad \downarrow \quad \downarrow \quad \downarrow$
 $\quad 2 \quad 1 \quad 1 \quad 0$
 Simple function

$$E[X] = 2 \cdot \frac{1}{4} + 1 \cdot \frac{1}{4} + 1 \cdot \frac{1}{4} + 0 \cdot \frac{1}{4} = \textcircled{1}$$

Prop $E[X]$ is uniquely determined by the distribution \mathcal{L}_X of X

1. For simple r.v's $X = a_i$ with prob. p_i , i.e.

$$\text{If } X = \sum_{i=1}^n a_i \mathbb{1}_{E_i} \Rightarrow E[X] = \sum_{i=1}^n a_i P(E_i) \stackrel{\text{def of } \int}{=} \sum_{i=1}^n a_i P\{X_i = a_i\} = \sum_{i=1}^n a_i \mathcal{L}_X(\{a_i\}). \checkmark$$

2. General case is determined by simple r.v's

Expressing $\mathbb{E}X$ in terms of distribution \mathcal{L}_X ?

Note: R.v. X on $(\mathbb{R}, \mathcal{F}, \mu)$ has the same distr. as r.v. $Y(x) := x$ on $(\mathbb{R}, \mathcal{B}, \mathcal{L}_X)$

$$\forall B \in \mathcal{B}: \underbrace{P\{X \in B\}}_{\mathcal{L}_X(B)} \stackrel{?}{=} \underbrace{P\{Y \in B\}}_{\mathcal{L}_X(\{x \in \mathbb{R} : Y(x) \in B\})}$$

Prop p. 18

$$\Rightarrow \mathbb{E}X = \mathbb{E}Y = \int_{-\infty}^{\infty} x d\mathcal{L}_X \quad \Downarrow$$

$$\mathcal{L}_X(\{x \in \mathbb{R} : x \in B\}) = \mathcal{L}_X(B) \quad \square$$

Prop \forall r.v. X ,

$$\mathbb{E}X = \int_{-\infty}^{\infty} x d\mathcal{L}_X$$

More generally: \forall mble $g: (\mathbb{R}, \mathcal{B}) \rightarrow (\mathbb{R}, \mathcal{B})$:

$$\mathbb{E}g(X) = \int_{-\infty}^{\infty} g(x) d\mathcal{L}_X$$

(same proof; note that $X \stackrel{\text{dist}}{=} Y \Rightarrow g(X) \stackrel{\text{dist}}{=} g(Y)$)

Expressing $\mathbb{E}X$ in terms of CDF: $F(x) = P\{X \leq x\} = \mathcal{L}_X((-\infty, x])$:

$$\mathbb{E}X = \int_{-\infty}^{\infty} x dF(x) \quad \leftarrow (\text{def of Lebesgue-Stieltjes integral})$$

PROPERTIES OF EXPECTATION (follow from properties of Lebesgue integral):

• Linearity: $\mathbb{E}\left[\sum_{i=1}^n a_i X_i\right] = \sum_{i=1}^n a_i \mathbb{E}X_i$

• Constant: $X = a$ a.s. $\Rightarrow \mathbb{E}X = a$.

• Monotonicity: $X \leq Y$ a.s. $\Rightarrow \mathbb{E}X \leq \mathbb{E}Y$

• Dominated Convergence: $X_n \rightarrow X$ a.s. $\&$ $|X_n| \leq Y$ a.s. $\forall n$ $\&$ $\mathbb{E}Y < \infty$
 $\Rightarrow \mathbb{E}X_n \rightarrow \mathbb{E}X$

Examples: (b) Σ angles of \forall triangle $= \pi$

A probabilistic proof:

• Project the Δ onto a random line (uniform orientation):

• # disappearing vertices = 1 always

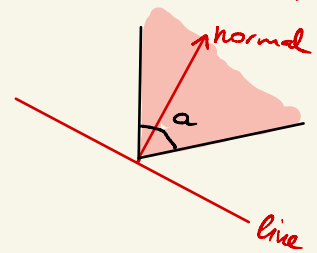
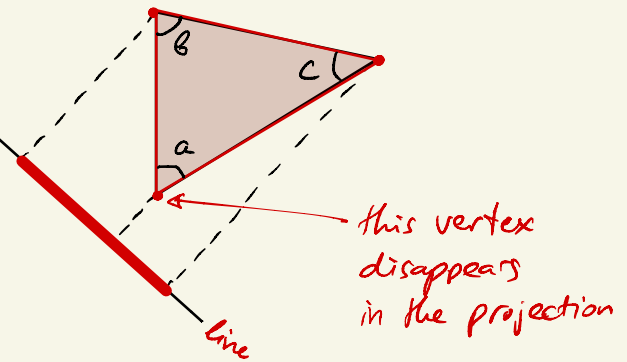
• $X_a := \begin{cases} 1 & \text{if vertex a disappears} \\ 0 & \text{otherwise} \end{cases}$; X_b, X_c similarly

$$\Rightarrow X_a + X_b + X_c = 1 \quad \text{a.s.}$$

$$\Rightarrow E[X_a] + E[X_b] + E[X_c] = 1$$

• But $E[X_a] = P\{\text{vertex a disappears}\}$ (def of Lebesgue integral: $\int \mathbb{1}_E dP = P(E)$)
 $= P\{\text{normal of the line} \in \text{cone}(a)\} = \frac{a}{\pi}$

$$\Rightarrow \frac{a}{\pi} + \frac{b}{\pi} + \frac{c}{\pi} = 1 \quad \Rightarrow \boxed{a + b + c = \pi}$$



(b) (Derangements p. 11)

What is the expected # of students who receive their own exam?

!!
X

$X = \sum_{i=1}^n X_i$ where $X_i = \begin{cases} 1 & \text{if student } i \text{ receives own exam} \\ 0 & \text{if not} \end{cases}$

$$\Rightarrow EX = \sum_{i=1}^n EX_i = \sum_{i=1}^n \underbrace{P\{\text{student } i \text{ receives own exam}\}}_{1/n} = 1.$$

1/n

DISCRETE DISTRIBUTIONS

Def X has discrete distr. if \mathcal{L}_X is supported on a finite or countable set

\uparrow
 X takes on finite or countable # of values x_1, x_2, \dots

\Rightarrow Probability mass function $x_i \mapsto P\{X=x_i\}$ determines distr of X .

$$EX = \int_{-\infty}^{\infty} x d\mathcal{L}_X = \sum_i x_i P\{X=x_i\} \quad \text{More generally: } \mathbb{E}g(x) = \sum_i g(x_i) P\{X=x_i\}$$

Ex

(a) Bernoulli distribution: $X \sim \text{Ber}(p)$ if $P\{X=1\}=p, P\{X=0\}=1-p$.

$$EX = 1 \cdot p + 0 \cdot (1-p) = p$$

(b) eg. indicator $X = \mathbb{1}_E$ of an event E . $X \sim \text{Ber}(p)$ with $p = P(E)$.

(c) $X = \# \text{Heads in 2 flips}$

$$EX = 0 \cdot P\{X=0\} + 1 \cdot P\{X=1\} + 2 \cdot P\{X=2\} = 0 \cdot \frac{1}{4} + 1 \cdot \frac{1}{2} + 2 \cdot \frac{1}{4} = 1$$

(d) $X \sim \text{Unif}\{x_1, \dots, x_n\} \Rightarrow P\{X=x_i\} = \frac{1}{n} \Rightarrow EX = \frac{1}{n} \sum_i x_i = \text{arithmetic mean}$

(e) (St. Petersburg Paradox)

The initial stake (money in the pot) is \$2.

\$2

→ The player tosses a fair coin.

If T \Rightarrow game ends, the player wins whatever is in the pot.

If H \Rightarrow stake doubles, game continues.

What would be a fair price to pay the casino for entering this game?

$X = \text{player's winnings}$.

$$P\{X=2\} = \frac{1}{2}; P\{X=4\} = \frac{1}{4}, \dots$$

$$E[X] = 2 \cdot \frac{1}{2} + 4 \cdot \frac{1}{4} + 8 \cdot \frac{1}{8} + \dots = \infty$$

Tosses	Winnings \$	Prob.
T	2	$\frac{1}{2}$
HT	4	$\frac{1}{4}$
HHT	8	$\frac{1}{8}$
HHT	16	$\frac{1}{16}$
...

} $\frac{1}{2}$ ← Median
 $\frac{1}{2}$

Ans: ∅ price.

Remarks: • Volatility, financial bubbles.

• Median is more stable. $\text{Med}(X) = M$ if

$$P\{X \leq M\} \geq \frac{1}{2}, P\{X \geq M\} \geq \frac{1}{2}$$

• In the St. Peter's paradox, $\text{Med}(X) \in [2, 4]$

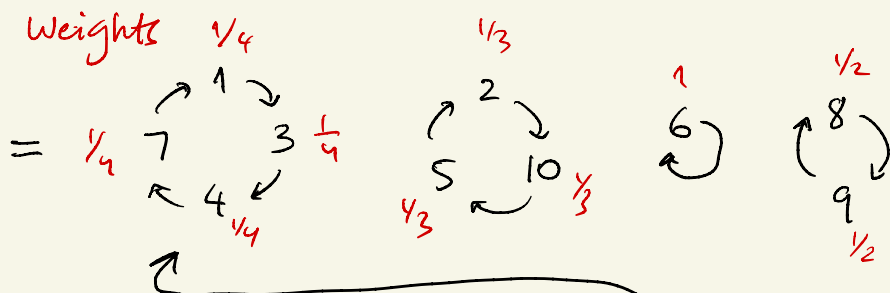
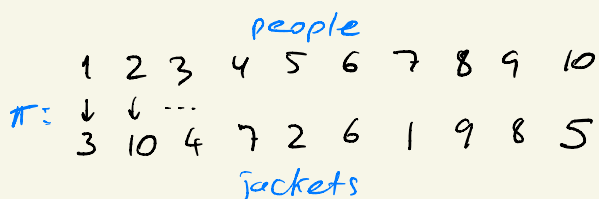
(f): Number of cycles in a random permutation

• Recall: permutation = a rearrangement of n distinct objects.

There are $n!$ permutations

• \forall permutation breaks down into cycles.

For example:



Q) What is $E \#(\text{cycles})$ in a random permutation?

↑
drawn from the set of all $n!$ permutations with uniform probability.

• Intuition: 1st cycle has length $\sim n$ (e.g. $n/2$), next $n/4, \dots \Rightarrow \log n$.

• Thm The expected # of cycles in a random permutation of n elements is ~~$\approx \ln n$~~

$$\sum_{k=1}^n \frac{1}{k} \approx \ln n + 0.58$$

Proof • To each element in a cycle of length k , assign weight $:= \frac{1}{k}$.

Sum of all weights = # cycles

E.g, here sum = $1+1+1+1=4$

• Formally, let

$Y_j :=$ length of the cycle the element j is in;

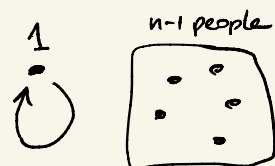
$$X = \sum_{j=1}^n \frac{1}{Y_j}$$

linearity $\Rightarrow E[X] = \sum_{j=1}^n \underbrace{E\left[\frac{1}{Y_j}\right]}_{\substack{\uparrow \\ \text{all equal by symmetry}}} = n \cdot \underbrace{E\left[\frac{1}{Y_1}\right]}_{\text{?}}$

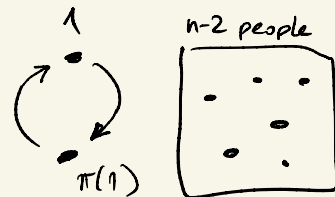
$$E\left[\frac{1}{Y_1}\right] = \sum_{k=1}^n \frac{1}{k} \underbrace{P\{Y_1 = k\}}_{\text{?}}$$

$P\{\text{element 1 is in a cycle of length } k\} = ?$

• $P\{1 \text{ is in a cycle of length } 1\} = \frac{(n-1)!}{n!} = \frac{1}{n}$ ← # ways to permute other $n-1$ people



$P\{1 \text{ is in a cycle of length } 2\} = \frac{(n-1)(n-2)!}{n!} = \frac{1}{n}$ # ways to choose $\pi(1)$ # ways to permute other $n-2$ people



...
 $P\{1 \text{ is in a cycle of length } k\} = \frac{1}{n}$ (DIT)

• $\Rightarrow E\left[\frac{1}{Y_1}\right] = \sum_{k=1}^n \frac{1}{k} \cdot \frac{1}{n} \Rightarrow E[X] = \sum_{k=1}^n \frac{1}{k}$ QED